The Effect of Child Gender on Parents' Labor Supply: Responses among Natives, Immigrants, and Racial and Ethnic Subgroups

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Abstract:

This paper examines whether the differential labor supply response of mothers and fathers by child gender varies between immigrants and natives and over racial and ethnic subgroups using the 1990–2000 Census and 1994–2008 March CPS supplement. We find that immigrants worked fewer weeks and hours per year if they have a son rather than a daughter. However, even the effect of having a son versus a daughter varies by whether or not the parents emigrated from an Asian country or another part of the world. We also find evidence in the CPS that the effect of child gender on men's labor supply is different for different racial groups, suggesting that son preference may be persistent even when the economic rationale for son preference no longer exists.

Keywords: child gender, son preference, immigrant, culture, Asian
JEL codes: J61, J16, J15, J13

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I. Introduction

There has long existed a tradition of son preference in many Asian countries. Recently, research on U.S. parents has shown that child gender affects many aspects of parents' behavior, including parents' labor supply (Lundberg and Rose 2002, Lundberg 2005), marital stability (Dahl and Moretti 2004, Lundberg and Rose 2003), and time spent with children (Yeung, et al. 2001, Lundberg, Pabilonia, and Ward-Batts 2008). Using the Panel Study of Income Dynamics (PSID), Lundberg and Rose (2002) found that white fathers' labor supply and wage rates increase more in response to having sons versus daughters while Lundberg (2005) found that non-black, non-Hispanic fathers of young sons in the NLSY79 work less than fathers of young daughters if their wives are college-educated. If sons increase marital stability, then we would expect that parents will respond by increasing specialization in marriage, i.e. fathers working more and mothers working less, unless the wife earns more than her husband. One hypothesis for why there exists differential parental behavior by child gender is that parents prefer sons or fathers prefer sons. This seems like a reasonable explanation for differential parental behavior by child gender in some developing countries where sons provide old age support. However, it is unclear why son preference would be a significant explanation for differential parental behavior by child gender in the U.S., where Social Security and private pensions are used as old age support and there is greater gender equality in the workplace. Almond and Edlund (2008), however, document an increase in the male/female ratio among U.S. second- and later-born children to parents of Asian descent. Another hypothesis for differential behavior is differences in the productivity of parental inputs in the production of children — either that sons require more time or fathers' time is more important in raising sons. If the latter is true, then fathers may spend more time with sons and less time in the labor market. It is also possible that fathers enjoy

spending time with sons, and then consequently will work less. Therefore, the magnitude and direction of any effect of child gender on labor market outcomes is an empirical question.

Hiedemann and Joesch (2005) provide an excellent review of the child gender literature through the lens of race and ethnicity. They find different effects of child gender on divorce, fertility and non-relative childcare by race[1], but do not find any evidence in the literature of differences by race in fathers' labor market outcomes. They conclude that more research on differences by race is necessary.

Using the 1994-2006 March Current Population (CPS) Supplements, Pabilonia and Ward-Batts (2007) examined whether having a son had differential effects on parent labor supply depending upon whether the parent was native born or a first-generation immigrant, and specifically whether the parent emigrated from an Asian country. We found some differential behavior among immigrants relative to natives, but even stronger evidence of differential effects by race. Specifically, Asian men and "other race" men worked less relative to white men in response to having a male child relative to a female child. This paper extends that analysis by examining the differential labor supply response of mothers and fathers by child gender using several large nationally representative samples. We will also examine whether the differing results from the PSID and NLSY79, found by Lundberg and Rose (2002) and Lundberg (2005), are due to sample composition in a later draft of this paper. In the present draft of this paper, we use the 5% sample of the 1990 and 2000 Census,[2] and the 1994-2008 March CPS. Results from Census data may differ due to changes over time in the pool of immigrants, the length of time immigrants have been in the U.S., and changes in the labor supply behavior in the immigrant's

[1] For example, Hiedemann and Joesch, (2005) find that infant and toddler sons are more likely than daughters to attend non-relative care if their mothers are black.
[2] Births in the NLSY79 occurred mostly in the 1980s so the sample should be more similar to the 1990 Census, although the 1990 Census will still include a newer wave of immigrants.

country of origin. Another possible difference between groups in the 1990 versus 2000 Census data is that their race may be classified differently due to the ability to indicate multiple races in 2000, but not in 1990.[3] Given the significant findings for Asians, regardless of nativity, in Pabilonia and Ward-Batts (2007), we also herein further our examination of immigrants in the CPS by looking at labor supply effects by child gender for second-generation immigrants. If son preference is the only explanation for differences in parental behavior, then immigrant status should magnify any effects of having a son on parents' labor supply. We would expect that any son preference effect would be smaller among second- and later-generation immigrants who have had more time to assimilate than among first-generation immigrants. If culture is persistent, then a tradition of son preference may persist even when the economic factors that generated such a preference no longer apply. If this is the case, we would expect to find similar effects of child gender within immigrant groups, regardless of country of birth, but differences across these groups. However, if differential parental labor supply behavior by child gender is due to differences in child production functions, then we expect that having sons rather than daughters would have a similar effect upon the parents' labor supply for both U.S.-born and immigrant parents, given the greater likelihood of similar adult outcomes in the U.S. than in developing countries.

II. Data and Methodology

[3] We classify those indicating two or more races in 2000 as mixed race, a category we do not have in 1990. 3.6% of immigrants in our sample fall into this race category, while less than 1% of the native-born sample does.

We use both the 1990 and 2000 Census five percent samples and the 1994-2008 March CPS Supplements.[4] In a future draft of the paper, we will include the 1980 Census. We examine the civilian labor supply behavior of married individuals aged 20-64 with only one child under age three in single-family households.[5] Therefore, the mothers and fathers that we examine are from the same households; however, we analyze their behavior separately.[6] We exclude families where one or both spouses is/are a farm worker, lives on a farm or in group housing, or is/are enrolled in school.[7]

We estimate reduced-form labor supply regressions using three different labor supply measures: hours worked last week (actual hours for 1990 Census and CPS, but usual hours for 2000 Census), weeks worked last year, and annual hours worked last year. Annual hours worked last calendar year are weeks worked last year multiplied by usual hours worked each week last year. For weeks worked last year and annual hours last year in all of our samples, as well as for usual weekly hours last year in the 2000 Census data, we examine the behavior in the previous calendar year of parents who currently have a child aged one or two. By excluding parents with infants, we avoid analyzing parental labor supply that occurred (at least partly) prior to the birth of the child.

Our main variables of interest are the gender of the parent's first (and only) child and child gender interacted with different measures of the respondent's immigrant status, and then

[4] The analysis for the CPS is for 1994 and after because information on country of birth was collected beginning in 1994. The Census data was taken from IPUMS files, which are located at http://usa.ipums.org/usa/, and the CPS data was from Unicon.

[5] We exclude younger parents aged 18 and 19 to exclude those who have a teenage pregnancy.

[6] Much of the literature on married couple labor supply assumes a secondary earner model. We don't do that here in that we do not assume the wife takes the husband's earnings as given. We include only household unearned income for both spouses. We do not have convincing exclusion restrictions available in these data that would allow us to estimate a joint model. In a later draft, we can include information on whether our results for wives would differ if we imposed the secondary earner model.

[7] In the CPS, we exclude the following 2002 Census occupation codes: 6040 and 6050. In the Census, we exclude the following 1990 Census occupation codes: 479, 483, 484, and 488.

alternatively, interacted with racial and Hispanic ethnicity group. Immigrant status is defined as being born outside of the 50 U.S. States. Thus, we include those born in U.S. territories, such as Puerto Rico and Guam, as immigrants. The respondent's degree of assimilation may correspond to whether they attended high school in the U.S. or the length of time spent in the U.S. Therefore, we also estimate specifications that include measures for whether the respondent moved to the U.S. before or after turning age 18.[8] We also allow for the possibility that Asian immigrants have a different effect of son compared to other groups. We define Asian immigrants as those who were born in Asia, including both southeast and southwest Asia.[9] We control for race and ethnicity using five mutually-exclusive categorical variables: white, black, Asian including Hawaiian/Pacific Islander, Native American/Alaskan/other race, and Hispanic.[10] In the 1996-2002 CPS, the Census recoded "other race" variables across the four race variables. Thus, there is likely measurement error in our race variables, which would bias our differential effects for Asian if it is not randomly assigned. Thus, when we look at differential racial effects, we also estimate a specification on data from the 2003-2008 CPS where "other race" can be grouped separately from Asian.

In addition, each model includes the usual demographic and human capital controls: age (and its square), spouse's age (and its square), family non-labor income, dummy variables for age of child, own education, spouse's education, geographic region, and, for CPS only, year in the sample (1990 and 2000 Census are analyzed separately). Family non-labor income in the CPS is converted to 1984-based real dollars using the CPI-U. All analyses are weighted. We

[8] Because respondents in some years were allowed to report an interval over which they entered the U.S. (varying from 1 to 10 years depending upon the survey year), we also include an indicator for the age of entry is ambiguous and set the other age of entry variables at zero. For the 2000 Census, respondents specified the exact year of entry.
[9] These countries include Bangladesh, Myanmar, Cambodia, China, Hong Kong, India, Indonesia, Japan, Korea, South Korea, Laos, Malaysia, Pakistan, Philippines, Singapore, Taiwan, Thailand, Vietnam, and Asia, not specified.
[10] If an individual is Hispanic, their race categorical variables are coded as 0.

report standard errors for the CPS regression analyses that correct for households that are sampled in consecutive March samples.

The sample size for the 1990 Census for the hours worked last week sample includes 95,365 mothers and 95,365 fathers. For hours last year and weeks last year, the sample includes 64,442 mothers and 64,442 fathers. Due to a lower birth rate in the late 1990s than in the late 1980s, the our sample of parents with a young child in the 2000 Census is slightly smaller, even though the population grew over that decade (see Statistical Abstract for the U.S. 2008). In the 2000 Census data, we are using labor supply variables solely from the previous year. For that sample, we have 54,120 mothers and 54,120 fathers. For the pooled 1994-2008 CPS sample, we have 19,441 mothers and 19,411 fathers in the actual hours worked last week sample. For the hours worked last year and weeks worked last year in the CPS samples, there are 12,835 mothers and 12,835 fathers.

Table 1 reports summary statistics for the hours worked last week sample and for the mean of weeks worked last year and annual hours where the sample size is reduced. On average, women with young children work about half as much as men with young children. Fathers are about 2 years older on average than mothers. Mothers and fathers have similar distributions across race and ethnicity categories as well as across educational attainment categories. Between 1990 and 2000, there is a six percent increase in the number of first-generation immigrants with one child under the age of three, which is also a six percentage point increase in our sample of parents comprised of this group.[11] Fifty-one to fifty-two percent of first-born children are boys, which is not statistically different from an expected child gender ratio of 1.05 males for every female if child gender is random. In addition, we did not find that the child gender ratio for

[11] First-generation Hispanic immigrants also increased from about 3.4% to 6.6% of our sample of parents between 1990 and 2000 Censuses.

immigrant parents or Asian immigrant parents was significantly different than the expected male to female ratio. Therefore, we conclude that there is no significant sex selection occurring in these samples for first births and the gender of the first-born child can be treated as exogenous in our regression analysis.

III. Results

A. Effects of Child Gender on Labor Supply for Immigrants Versus U.S.-born

Table 2 presents OLS regression estimates of weeks worked per year, annual hours worked, and hours worked last week in the 1990 Census for fathers and mothers respectively. In the first specification, we do not allow for differences between natives and immigrants. In the second specification, we allow for differences between U.S.-born and first-generation immigrants, both in levels of the outcome and in the response to having a son rather than a daughter. In the third specification, we allow for immigrants who entered the U.S. at different ages to differ both in the level of the outcome, as well as in their response to having a son rather than a daughter. In 1990 Census data, the year of immigration is given as a range of years. Depending on the specific range of years and current age, we know whether most immigrants arrived in the U.S. prior to reaching age 18 versus at age 18 or older. However, for some, we do not know whether they were yet age 18 at the time of entry, but just that they were within a few years of age 18 (within 5 years for most of these) at the time of entry. Therefore, we also include a separate category for those whose age at entry is ambiguous, as well as an interaction of that variable with son. Finally, in the fourth specification, we allow for differences between Asian immigrants and other immigrants. Similarly, Table 3 presents results using the 2000 Census.[12]

[12] A full set of regression estimates are available upon request.

In 2000 Census data, the specific year of immigration is known, so the ambiguous age at entry category is not required in estimates for 2000.

We find that fathers worked more weeks if they had a son than a daughter in 1990 but the magnitude of the effect is very small. In 1990, when we allow for immigrants who entered at different ages to have different effects, we find that both immigrant mothers and fathers with a son worked fewer weeks relative to immigrant parents with a daughter. The estimates are a bit larger, at just under one week for fathers and just over 2 weeks for mothers. A joint test of this coefficient with that for son is also significant, indicating that immigrant mothers and fathers work fewer weeks in response to having a son rather than a daughter than is the case for natives. However, this effect is undone for immigrants who entered the US as adults – a joint test of the three son and son interaction coefficients is not statistically different from zero. In contrast to 1990, we find no statistically significant effects of having a son relative to a daughter in 2000.

In 1990, we find that immigrant fathers worked fewer hours in the last year if they had a son compared to a daughter. However, immigrants who entered the U.S. at or after age 18 have a significant effect for son which offset this decrease in hours, while Asian immigrant fathers have a similar positive effect for son which more than offsets the negative son effect associated with immigrants on average. However, the positive net effect of son for Asian immigrants is not statistically significant. Again, when we allow for age of entry effects, we find that immigrant mothers with a son worked fewer hours the previous year than did those with a daughter, but this result does not hold for immigrants who entered the US as an adult.

We do not find any significant differential effects of a son on parents' annual hours worked in the 2000 Census data. We do find that the usual weekly hours reported for the previous year are significantly less for fathers with a son relative to those with a daughter in

2000. However, we do not find any significant differences in the response to child gender between natives and immigrants for this variable in 2000.

In Table 4, we repeat the analyses from Tables 2 and 3, but omit couples in which one partner is an immigrant and the other is a native. This reduces our sample size slightly to 60,207 for weeks and annual hours in 1990, 89,075 for hours last week in 1990, and 49,524 for all labor supply variables in 2000. Generally we find similar estimates to those from Tables 2 and 3, with slight variations in magnitude and level of statistical significance. In particular, for annual hours in the 1990 data, we find larger son effects for immigrant fathers and for the offsetting effect for immigrants who entered as adults.

We present OLS regression estimates of weeks worked per year, annual hours worked, and hours worked last week in the CPS for fathers and mothers in Tables 5-7, respectively.[13] In the first specification for each outcome, we do not allow for differences between U.S.-born and immigrant parents. In the second specification, we allow for differences between those born in the U.S. and first-generation immigrants, both in levels of the outcome and in the response to having a son rather than a daughter. In the third specification, we allow for different immigrant effects based upon age of entry in the U.S (equal to or greater than age 18). In the fourth specification, we allow for differences among U.S.-born, first-generation immigrants and second-generation immigrants. In the fifth specification, we allow for differences between Asian immigrants and other immigrants. In the final specification, we allow for differences among U.S.-born, first- and second-generation Asian immigrants, and first- and second-generation non-Asian immigrants.

In the CPS, we do not find any differences in parents' labor supply behavior by child gender for those parents born in the U.S. First-generation immigrant men with a son rather than

[13] We also estimated censored regression models for mothers. Results are similar and thus not reported.

a daughter work over one week less and about 65 fewer hours annually. This effect holds for all of the specifications. In column 3 (Tables 5-7), we find that the strong negative effect of having a son versus a daughter on father's labor supply, relative to U.S.-born fathers, is for those who emigrated after age seventeen. (This is in contrast to our findings from Census data. This may be attributable to a different composition of immigrants and we will investigate this difference further.) Even though the individual coefficient estimates on the interaction effect of immigrant age greater than or equal to age 18 with son is not significant, it is jointly significant with the interaction effect of immigrant and son. In a sensitivity analysis, we exclude those parents who are part of an immigrant–U.S.-born couple. The effects are similar (Table 8). We also estimated a specification where we defined immigrant as being a first- or a second-generation immigrant who is also married to a first- or second-generation immigrant (see Appendix Table 1 for results) and we excluded those immigrants who are not married to a first- or second-generation immigrant from the sample. We find strong negative effects of having a son rather than a daughter on weeks worked and annual hours worked last year for immigrant fathers.

We also find that second-generation Asian immigrant fathers with a son rather than a daughter worked 9 more hours in the previous week. This is almost entirely offset by a decrease in second-generation Asian immigrant mother's work hours the previous week if they have a son rather than a daughter. These latter results suggest that sons increase the traditional gender specialization within Asian families – with fathers working more and mothers working less.

B. Effects of Child Gender on Labor Supply By Racial and Ethnic Groups

We next explore the possible differences between racial and ethnic groups within the U.S. since we have some evidence that a cultural preference for sons may be persistent. For each outcome, we first include a son dummy and interactions of this variable with race indicators for

Black, Asian, and other (non-white) race as well as for Hispanic ethnicity. We add an immigrant indicator in the second specification. Finally, we add an immigrant-son interaction in the third specification.

Tables 9 and 10 present results from 1990 and 2000 Census data respectively. We find significant effects of having a son rather than a daughter in this specification: fathers work slightly more weeks in the year in 1990 and fewer annual hours and usual weekly hours in 2000. We also find a negative significant effect on annual hours for immigrant fathers with a son relative to those with a daughter in 1990. The only significant differential son effects by race that we find are negative effects on hours last week among our "other race" category, which includes Native Americans and Alaskan Natives.

However, in CPS data we find that Asian mothers with a son rather than a daughter work 2 ½ to 4 fewer hours per week (Table 11). This is consistent with the previous result for second-generation Asian immigrant women. We also find that immigrant mothers with a son rather than a daughter work a few more hours per week (last week), regardless of race. We find no significant child gender differences between racial or ethnic groups for mothers in either weeks or hours worked. However, we find that, compared to non-Hispanic white men, men in the Asian and "other race" category worked fewer hours last year if they have a son rather than a daughter. This result is robust to the addition of the immigrant dummy, but when the immigrant-son interaction is included, the result for men of other race remains significant while the result for Asian men is no longer significant at conventional levels. We also find a significant positive effect of having a son on annual hours for fathers (column 6), which is consistent with the findings of Lundberg and Rose (2002). In addition, we find that "other race" fathers worked about 6 fewer hours last week if they have a son rather than a daughter compared to native-born

fathers. In Table 12, we estimate similar specifications on a restricted sample for years 2003-2008 to examine whether measurement error in our Asian variable in the preceding years may bias our results. Indeed, we find much stronger and significant negative effects on both weeks worked and annual hours worked for Asian fathers who have a son rather than a daughter. These differential effects across racial groups may explain some of the significant child gender effects other researchers (e.g. Lundberg 2005) have found for parental labor supply behavior.

IV. Conclusion

We find some evidence that there are differential child gender effects on parents' labor force behavior between immigrants and natives – immigrants work fewer weeks and hours per year if they have a son rather than a daughter. However, even the effect of having a son versus a daughter varies by whether or not the parents emigrated from an Asian country or another part of the world. We also find evidence in the CPS that the effect of child gender on men's labor supply is different for different racial groups, suggesting that son preference may be persistent even when the economic rationale for son preference no longer exists. Asian men and particularly men in the "other race" category work less relative to non-Hispanic white men if they have a son rather than a daughter. This suggests that there may be a decline in specialization within marriage for these groups relative to non-Hispanic whites when the couple has a son rather than a daughter. It could be attributable to men's greater desire to spend more time with sons, especially young sons. However, it could also mean that sons require father's time more than daughters do.

We find some contrast in results from the Census data relative to those from the CPS. These differences require further investigation into possible differences in the composition of the

sample of parents, as well as variable definition differences. We hope to report further on these issues in a future draft.

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Table 1: Summary Statistics

	1990 Census				2000 Census				1994-2008 CPS			
	Household-level variables				Household-level variables				Household-level variables			
Variable	Mean		Std. Dev.		Mean		Std. Dev.		Mean		Std. Dev.	
Son	0.51				0.51				0.52			
Child age one	0.39				0.35				0.36			
Child age two	0.29				0.27				0.28			
Non-labor income	1,746		5,878		3,140		13,246		1533		5271	
	Mothers		Fathers		Mothers		Fathers		Mothers		Fathers	
	Mean	Std. Dev.	Mean	Std Dev	Mean	Std. Dev.	Mean	Std Dev	Mean	Std. Dev.	Mean	Std. Dev.
Hours last week (Census 2000 Usual hours)	20.49	19.38	43.35	15.40	28.03	20.15	44.53	12.18	20.73	24.67	41.05	20.69
Weeks last year	30.47	21.94	48.02	10.30	31.51	22.48	48.00	10.89	32.48	35.03	48.88	15.25
Hours last year	1098.43	910.40	2177.26	682.66	1135.53	1045.94	2202.11	700.24	1179.04	1455.03	2188.43	991.77
First-generation immigrant	0.09		0.09		0.16		0.16		0.18		0.18	
US entry age >=18	0.04		0.04		0.10		0.10		0.11		0.11	
US entry age<18	0.03		0.03		0.06		0.06		0.06		0.06	
Asian immigrant	0.03		0.03		0.05		0.04		0.05		0.05	
Black	0.04		0.04		0.05		0.05		0.05		0.06	
Asian/Pacific Islander	0.03		0.03		0.06		0.05		0.06		0.05	
Other race	0.01		0.01		0.01		0.01		0.01		0.01	
Hispanic	0.07		0.06		0.12		0.12		0.11		0.11	
Age	28.38	5.01	30.86	5.87	30.07	5.65	32.40	6.39	29.85	7.06	32.23	8.15
Education												
High School	0.33		0.33		0.32		0.36		0.23		0.25	
Some college	0.33		0.29		0.24		0.22		0.27		0.26	
College Degree	0.29		0.31		0.38		0.35		0.31		0.28	
Advanced Degree									0.14		0.14	
N	95,365		95,365		54,120		54,120		19,441		19,441	

Note: Survey weights used. Family non-labor income in the CPS is reported in 1984 dollars. Non-labor income in Census data is in nominal dollars.

Table 2. Differential Effects of Child Gender on Parental Labor Supply for Immigrants versus U.S.-born (1990 Census)

A. Weeks Last Year	Fathers				Mothers			
Son	0.15*	0.19**	0.19**	0.19**	-0.06	0.05	0.04	0.05
	(0.09)	(0.09)	(0.09)	(0.09)	(0.18)	(0.19)	(0.19)	(0.19)
Immigrant*son		-0.39	-0.94*	-0.72		-0.99	-2.20**	-1.07
		(0.37)	(0.52)	(0.44)		(0.66)	(1.02)	(0.77)
US entry age ≥18*son			1.18				1.79	
			(0.77)				(1.37)	
Asian immigrant *son				1.26				0.26
				(0.78)				(1.41)
Immigrant		-0.77***	0.80**	-0.57*		-4.26***	1.37*	-4.14***
		(0.30)	(0.36)	(0.33)		(0.53)	(0.75)	(0.58)
US entry age ≥18			-3.53***				-11.71***	
			(0.57)				(1.00)	
Asian immigrant				-1.04				-0.79
				(0.72)				(1.34)
R-squared	0.05	0.05	0.06	0.05	0.06	0.07	0.07	0.07

B. Annual Hours Last Year	Fathers				Mothers			
Son	0.68	5.42	5.41	5.41	0.30	4.16	4.15	4.18
	(5.75)	(5.89)	(5.89)	(5.89)	(7.65)	(7.97)	(7.97)	(7.97)
Immigrant*son		-47.73**	-106.13***	-69.48***		-36.17	-89.60**	-37.31
		(23.01)	(36.62)	(26.68)		(27.62)	(42.28)	(31.23)
US entry age ≥18*son			103.91**				89.78	
			(48.88)				(57.09)	
Asian immigrant *son				84.44*				3.29
				(50.01)				(60.69)
Immigrant		-13.43	85.45***	0.72		-135.28***	75.86**	-139.52***
		(18.76)	(28.24)	(20.61)		(22.17)	(31.63)	(23.86)
US entry age ≥18			-212.07***				-447.06***	
			(37.72)				(41.49)	
Asian immigrant				-73.18				39.08
				(46.76)				(58.34)
R-squared	0.05	0.05	0.05	0.05	0.06	0.06	0.06	0.06

Table 2. Cont'd: Differential Effects of Child Gender on Parental Labor Supply for Immigrants versus U.S.-born (1990 Census)

C. Hours Last Week					Mothers			
Son	-0.13	-0.13	-0.13	-0.13	-0.00	-0.01	-0.01	-0.01
	(0.11)	(0.11)	(0.11)	(0.11)	(0.13)	(0.14)	(0.14)	(0.14)
Immigrant*son		-0.05	-0.31	-0.35		0.06	-0.21	0.07
		(0.39)	(0.61)	(0.46)		(0.48)	(0.76)	(0.54)
US entry age ≥18*son			-0.01				0.31	
			(0.83)				(1.00)	
Asian immigrant *son				1.18				-0.06
				(0.84)				(1.03)
Immigrant		0.15	0.85*	0.33		-2.87***	0.42	-3.00***
		(0.32)	(0.46)	(0.35)		(0.38)	(0.56)	(0.41)
US entry age ≥18			-1.21**				-7.15***	
			(0.61)				(0.72)	
Asian immigrant				-0.93				1.06
				(0.80)				(1.00)
R-squared	0.04	0.04	0.04	0.04	0.06	0.06	0.07	0.06

Notes: Sample sizes are 64,442 for weeks and annual hours and 95,365 for hours worked last week. Standard errors are in parentheses. Survey weights are used. Significance levels: ** = 5%; * = 10%. Control variables include race, Hispanic ethnicity, age of child, quadratic in age of respondent and spouse, family non-labor income, region, education categories for respondent and spouse, and a constant. In the specifications with age of entry, we also control for ambiguous age of entry and its interaction with son.

Table 3. Differential Effects of Child Gender on Parental Labor Supply for Immigrants versus U.S.-born (2000 Census)

A. Weeks Last Year	Fathers				Mothers			
Son	-0.02	0.01	0.02	0.01	0.00	0.09	0.10	0.09
	(0.10)	(0.10)	(0.10)	(0.10)	(0.21)	(0.22)	(0.22)	(0.22)
Immigrant*son		-0.22	0.39	0.09		-0.46	0.08	-0.61
		(0.34)	(0.49)	(0.41)		(0.58)	(0.91)	(0.68)
US entry age ≥18*son			-0.90				-0.84	
			(0.64)				(1.10)	
Asian immigrant *son				-1.07				0.47
				(0.67)				(1.16)
Immigrant		-1.20***	-0.07	-1.34***		-7.86***	-2.87***	-7.67***
		(0.28)	(0.36)	(0.31)		(0.49)	(0.68)	(0.53)
US entry age ≥18			-1.87***				-8.20***	
			(0.45)				(0.79)	
Asian immigrant				0.32				-1.22
				(0.70)				(1.21)
R-squared	0.05	0.05	0.05	0.05	0.07	0.08	0.09	0.08
B. Annual Hours Last Year								
	Fathers				Mothers			
Son	-7.37	-10.38	-10.37	-10.45	-3.68	-2.26	-2.17	-2.28
	(6.48)	(6.81)	(6.81)	(6.81)	(8.69)	(9.39)	(9.39)	(9.39)
Immigrant*son		19.86	28.06	32.59		-5.54	-3.51	-19.34
		(20.15)	(29.51)	(23.83)		(24.08)	(37.92)	(28.01)
US entry age ≥18*son			-10.65				-3.04	
			(37.88)				(45.74)	
Asian immigrant *son				-43.07				43.39
				(40.83)				(48.45)
Immigrant		-83.03***	-2.96	-80.30***		-288.31***	-89.50***	-276.00***
		(16.79)	(22.40)	(18.36)		(20.19)	(28.50)	(21.98)
US entry age ≥18			-131.99***				-326.15***	
			(27.07)				(32.87)	
Asian immigrant				-72.32*				-66.77
				(43.89)				(51.08)
R-squared	0.05	0.05	0.05	0.05	0.06	0.07	0.07	0.07

Table 3. Cont'd: Differential Effects of Child Gender on Parental Labor Supply for Immigrants versus U.S.-born (2000 Census)

C. Usual Weekly Hours Last Year

	Fathers				Mothers			
Son	-0.21*	-0.26**	-0.26**	-0.26**	-0.06	-0.00	-0.00	-0.00
	(0.11)	(0.12)	(0.12)	(0.12)	(0.17)	(0.18)	(0.18)	(0.18)
Immigrant*son		0.35	0.31	0.49		-0.25	-0.31	-0.66
		(0.35)	(0.52)	(0.42)		(0.50)	(0.76)	(0.58)
US entry age ≥18*son			0.08				0.09	
			(0.67)				(0.94)	
Asian immigrant *son				-0.47				1.28
				(0.71)				(1.01)
Immigrant		-0.76**	0.29	-0.70**		-5.88***	-1.50***	-5.61***
		(0.30)	(0.40)	(0.33)		(0.41)	(0.57)	(0.45)
US entry age ≥18			-1.71***				-7.18***	
			(0.48)				(0.68)	
Asian immigrant				-1.09				-1.14
				(0.78)				(1.02)
R-squared	0.03	0.03	0.03	0.03	0.05	0.06	0.06	0.06

Notes: Sample sizes are 54,120 for weeks, annual hours and usual weekly hours last year. Standard errors are in parentheses. Survey weights are used. Significance levels: ** = 5%; * = 10%. Control variables include race, Hispanic ethnicity, age of child, quadratic in age of respondent and spouse, family non-labor income, region, education categories for respondent and spouse, and a constant.

Table 4. Effects of Child Gender for Two-Immigrant Parent Families versus Native Parent Families (1990 & 2000Census)

A. 1990: Weeks Last Year

	Fathers			Mothers		
Son	0.19**	0.19**	0.19**	0.03	0.03	0.03
	(0.09)	(0.09)	(0.09)	(0.19)	(0.19)	(0.19)
Immigrant*son	-0.03	-1.50	-0.47	-0.89	-2.59*	-0.81
	(0.50)	(1.02)	(0.64)	(0.82)	(1.54)	(0.99)
US entry age ≥18*son		2.30*			2.27	
		(1.21)			(1.87)	
Asian immigrant *son			1.28			-0.22
			(0.98)			(1.71)
Immigrant	-1.47***	0.81	-1.12**	-5.90***	1.18	-5.77***
	(0.44)	(0.73)	(0.50)	(0.69)	(1.17)	(0.77)
US entry age ≥18		-4.09***			-11.93***	
		(0.89)			(1.38)	
Asian immigrant			-1.68*			-1.21
			(0.98)			(1.66)
R-squared	0.06	0.06	0.06	0.07	0.07	0.07

B. 1990: Annual Hours Last Year

	Fathers			Mothers		
Son	6.25	6.25	6.23	2.91	2.89	2.92
	(6.00)	(6.00)	(6.00)	(8.13)	(8.13)	(8.13)
Immigrant*son	-27.45	-183.89***	-51.59	-29.38	-116.13*	-30.00
	(29.44)	(62.02)	(36.44)	(34.41)	(63.15)	(40.01)
US entry age ≥18*son		222.47***			128.14*	
		(72.35)			(77.27)	
Asian immigrant *son			70.67			1.88
			(60.05)			(73.24)
Immigrant	-65.23**	101.17**	-48.00*	-172.48***	104.72**	-175.30***
	(25.75)	(50.53)	(29.09)	(28.91)	(49.21)	(31.45)
US entry age ≥18		-272.70***			-476.28***	
		(57.19)			(57.52)	
Asian immigrant			-75.94			22.64
			(59.48)			(72.47)
R-squared	0.05	0.05	0.05	0.06	0.06	0.06

Table 4. Effects of Child Gender for Two-Immigrant Parent Families versus Native Parent Families (1990 & 2000Census)

C. 1990:Hours Last Week

	Fathers			Mothers		
Son	-0.14	-0.14	-0.14	-0.02	-0.02	-0.02
	(0.11)	(0.11)	(0.11)	(0.14)	(0.14)	(0.14)
Immigrant*son	0.07	-0.69	-0.32	-0.12	-0.05	-0.50
	(0.49)	(1.01)	(0.61)	(0.59)	(1.15)	(0.70)
US entry age ≥18*son		0.82			-0.04	
		(1.19)			(1.38)	
Asian immigrant *son			1.11			1.10
			(0.99)			(1.26)
Immigrant	0.09	1.22	0.38	-3.49***	0.46	-3.41***
	(0.42)	(0.76)	(0.48)	(0.50)	(0.86)	(0.55)
US entry age ≥18		-1.69*			-7.15***	
		(0.87)			(1.00)	
Asian immigrant			-1.36			0.29
			(1.02)			(1.25)
R-squared	0.04	0.04	0.04	0.06	0.07	0.06

D. 2000:Weeks Last Year

	Fathers			Mothers		
Son	0.01	0.01	0.01	0.17	0.17	0.17
	(0.11)	(0.11)	(0.11)	(0.23)	(0.23)	(0.23)
Immigrant*son	-0.44	0.79	-0.17	-0.86	-1.09	-1.19
	(0.41)	(0.70)	(0.52)	(0.67)	(1.20)	(0.81)
US entry age ≥18*son		-1.68**			0.26	
		(0.84)			(1.39)	
Asian immigrant *son			-0.79			0.93
			(0.77)			(1.34)
Immigrant	-1.61***	-0.68	-1.74***	-9.74***	-3.82***	-9.56***
	(0.36)	(0.54)	(0.41)	(0.59)	(0.91)	(0.65)
US entry age ≥18		-1.31**			-8.27***	
		(0.59)			(1.00)	
Asian immigrant			0.28			-0.52
			(0.90)			(1.46)
R-squared	0.05	0.05	0.05	0.08	0.09	0.08

T able 4. Effects of Child Gender for Two-Immigrant Parent Families versus Native Parent Families (1990 & 2000Census)

E.2000:Annual Hours Last Year		Fathers			Mothers	
Son	-9.93	-9.93	-10.01	1.12	1.17	1.10
	(6.96)	(6.96)	(6.96)	(9.62)	(9.62)	(9.62)
Immigrant*son	13.72	41.10	20.49	-25.71	-69.39	-51.18
	(23.73)	(40.16)	(29.80)	(27.56)	(48.98)	(32.68)
US entry age ≥18*son		-34.86			58.30	
		(48.10)			(56.76)	
Asian immigrant *son			-17.64			72.82
			(46.22)			(55.19)
Immigrant	-110.03***	-26.81	-106.02***	-346.28***	-98.02***	-328.89***
	(21.32)	(32.30)	(23.77)	(24.36)	(37.74)	(26.65)
US entry age ≥18		-118.60***			-347.14***	
		(35.04)			(40.93)	
Asian immigrant			-66.62			-71.51
			(53.59)			(61.37)
R-squared	0.05	0.05	0.05	0.07	0.07	0.07
F. 2000:UsualWeekly Hours Last Year						
		Fathers			Mothers	
Son	-0.26**	-0.26**	-0.26**	0.06	0.06	0.06
	(0.12)	(0.12)	(0.12)	(0.19)	(0.19)	(0.19)
Immigrant*son	0.32	0.31	0.38	-0.68	-1.41	-1.25*
	(0.42)	(0.70)	(0.53)	(0.58)	(1.03)	(0.70)
US entry age ≥18*son		0.07			0.96	
		(0.85)			(1.20)	
Asian immigrant *son			-0.14			1.63
			(0.80)			(1.18)
Immigrant	-1.05***	0.21	-0.99**	-7.27***	-1.91**	-6.88***
	(0.37)	(0.56)	(0.42)	(0.51)	(0.78)	(0.56)
US entry age ≥18		-1.80***			-7.49***	
		(0.62)			(0.87)	
Asian immigrant			-0.82			-1.60
			(0.94)			(1.24)
R-squared	0.03	0.04	0.04	0.06	0.06	0.06

Notes: Census 1990 sample sizes are 60,207 for weeks and annual hours and 89,075 for hours worked last week. Census 2000 sample sizes are 49,524 for weeks, annual hours and usual weekly hours last year. Standard errors are in parentheses. Survey weights are used.

Significance levels: ** = 5%; * = 10%. Control variables include race, Hispanic ethnicity, age of child, quadratic in age of respondent and spouse, family non-labor income, region, education categories for respondent and spouse, and a constant. In the specifications with age of entry in 1990, we also control for ambiguous age of entry and its interaction with son.

Table 5. Differential Effects of Child Gender on Weeks Worked Last Year for Immigrants versus U.S.-born (1994-2008 CPS)

	Fathers						Mothers					
Son	0.08	0.29	0.29	0.32	0.30	0.31	0.31	0.32	0.33	0.29	0.33	0.34
	(0.21)	(0.23)	(0.23)	(0.23)	(0.23)	(0.23)	(0.48)	(0.52)	(0.52)	(0.53)	(0.52)	(0.53)
Immigrant*son		-1.06*	-0.05	-1.09*	-0.97	-0.93		-0.11	-0.16	-0.07	0.47	0.39
		(0.61)	(0.76)	(0.62)	(0.63)	(0.65)		(1.26)	(2.04)	(1.26)	(1.40)	(1.42)
US entry age ≥18*son			-1.49						0.06			
			(1.10)						(2.42)			
2nd gen. immigrant*son				-0.76		-0.90				0.98		1.52
				(1.09)		(1.03)				(2.36)		(2.34)
Asian immigrant *son					-0.38	-0.39					-1.92	-1.86
					(1.51)	(1.51)					(2.67)	(2.65)
2nd gen. Asian immigrant *son						1.97						-7.11
						(2.32)						(5.28)
Immigrant		-0.42	0.13	-0.36	-0.23	--0.33		-8.73***	-2.63*	-8.50***	-8.72***	-8.68***
		(0.49)	(0.62)	(0.49)	(0.51)	(0.52)		(1.05)	(1.54)	(1.06)	(1.13)	(1.15)
US entry age ≥18			-1.22						-10.06***			
			(0.79)						(1.73)			
2nd gen. immigrant				0.91		0.98				2.74		2.63
				(0.75)		(0.67)				(1.71)		(1.72)
Asian immigrant					-1.54	-1.61					-1.22	1.66
					(1.32)	(1.36)					(2.44)	(2.57)
2nd gen. Asian immigrant						-1.01						0.95
						(2.10)						(3.62)
R-squared	0.041	0.042	0.045	0.043	0.043	0.043	0.098	0.110	0.118	0.111	0.111	0.112

Notes: Sample sizes are 12,835 fathers and 12,835 mothers. Standard errors are in parentheses. Survey weights are used. Significance levels: ** = 5%; * = 10%. Control variables include race, Hispanic ethnicity, age of child, quadratic in age of respondent and spouse, family non-labor income, region, year, education categories for respondent and spouse, and a constant. In the specifications with age of entry, we also control for ambiguous age of entry and its interaction with son.

Table 6. Differential Effects of Child Gender on Annual Hours for Immigrants versus U.S.-born (1994-2008 CPS)

	Fathers						Mothers					
Son	6.66	19.46	19.48	19.01	20.36	20.33	25.47	29.61	29.78	27.83	29.69	30.06
	(13.85)	(15.28)	(15.28)	(15.61)	(15.28)	(15.63)	(19.85)	(21.88)	(21.87)	(22.42)	(21.88)	(22.46)
Immigrant*son		-64.84*	-21.29	-64.36*	-58.93	-53.07		-22.27	-55.54	-20.42	2.98	-2.43
		(35.95)	(49.77)	(36.08)	(38.61)	(39.96)		(50.62)	(82.29)	(50.85)	(55.61)	(56.64)
US entry age ≥18*son			-58.14						47.08			
			(65.38)						(97.23)			
2nd gen. immigrant*son				17.11		22.80				48.29		79.59
				(78.41)		(78.21)				(100.92)		(101.24)
Asian immigrant *son					-25.16	-28.30					-83.59	-80.20
					(80.85)	(80.71)					(107.26)	(107.08)
2nd gen. Asian immigrant *son						57.82						-340.12
						(184.82)						(227.61)
Immigrant		-52.56*	1.72	-49.68	-31.37	-35.23		-311.81***	-65.65	-308.57***	-315.13***	-317.72***
		(30.88)	(38.62)	(30.98)	(33.34)	(34.11)		(42.74)	(63.53)	(43.08)	(45.31)	(46.35)
US entry age ≥18			-109.27**						-405.70***			
			(48.79)						(70.77)			
2nd gen. immigrant				36.61		16.81				30.90		12.43
				(53.51)		(52.85)				(73.21)		(73.64)
Asian immigrant					-163.37**	-114.92					-27.49	-23.47
					(73.69)	(77.62)					(104.83)	(110.29)
2nd gen. Asian immigrant						155.30						163.28
						(129.13)						(168.04)
R-squared	0.049	0.051	0.053	0.051	0.052	0.053	0.090	0.100	0.106	0.100	0.100	0.101

Notes: Sample sizes are 12,835 fathers and 12,835 mothers. Standard errors are in parentheses. Survey weights are used. Significance levels: ** = 5%; * = 10%. Control variables include race, Hispanic ethnicity, age of child, quadratic in age of respondent and spouse, family non-labor income, region, year, education categories for respondent and spouse, and a constant. In the specifications with age of entry, we also control for ambiguous age of entry and its interaction with son.

Table 7. Differential Effects of Child Gender on Hours Worked Last Week for Immigrants versus U.S.-born (1994-2008 CPS)

	Fathers						Mothers					
Son	0.18	0.29	0.29	0.25	0.30	0.21	0.09	0.02	0.03	0.01	0.02	0.05
	(0.29)	(0.32)	(0.32)	(0.33)	(0.32)	(0.33)	(0.34)	(0.38)	(0.38)	(0.39)	(0.38)	(0.39)
Immigrant*son		-0.57	0.39	-0.52	-0.50	-0.36		0.44	0.96	0.45	0.99	0.88
		(0.75)	(1.16)	(0.75)	(0.83)	(0.86)		(0.87)	(1.46)	(0.87)	(0.94)	(0.97)
US entry age ≥18*son			-1.43						-0.70			
			(1.43)						(1.69)			
2nd gen. immigrant*son				1.13		0.50				0.31		1.09
				(1.56)		(1.57)				(1.80)		(1.82)
Asian immigrant *son					-0.20	-0.23					-1.93	-1.90
					(1.61)	(1.61)					(1.85)	(1.86)
2nd gen. Asian immigrant *son						9.35**						-7.11*
						(3.92)						(3.72)
Immigrant		-2.01***	-1.51	-2.10***	-1.84***	-2.00***		-5.68***	-1.42	-5.69***	-6.19***	-6.24***
		(0.65)	(0.94)	(0.65)	(0.70)	(0.72)		(0.74)	(1.11)	(0.74)	(0.79)	(0.81)
US entry age ≥18			-1.00						-7.00***			
			(1.10)						(1.24)			
2nd gen. immigrant				-1.54		-1.45				-0.32		-0.89
				(1.06)		(1.07)				(1.27)		(1.27)
Asian immigrant					-1.44	-0.94					2.65	3.00
					(1.50)	(1.55)					(1.86)	(1.95)
2nd gen. Asian immigrant						-0.92						4.82*
						(2.45)						(2.70)
R-squared	0.034	0.036	0.036	0.036	0.036	0.037	0.073	0.079	0.085	0.079	0.079	0.080

Notes: Sample sizes are 19,441 fathers and 19,441 mothers. Standard errors are in parentheses. Survey weights are used. Significance levels: ** = 5%; * = 10%. Control variables include race, Hispanic ethnicity, age of child, quadratic in age of respondent and spouse, family non-labor income, region, year, education categories for respondent and spouse, and a constant. In the specifications with age of entry, we also control for ambiguous age of entry and its interaction with son.

Table 8. Effects of Child Gender for Two-Immigrant Parent Families versus Native Parent Families[a] (1994-2008 CPS)

	Weeks worked last year			Annual hours		
Fathers	(1)	(2)	(3)	(4)	(5)	(6)
Son	0.24	0.24	0.29	19.95	19.97	20.90
	(0.23)	(0.23)	(0.23)	(15.47)	(15.47)	(15.50)
Immigrant*son	-1.01	0.45	-1.17*	-63.02*	-26.98	-60.05
	(0.66)	(0.84)	(0.68)	(37.53)	(53.23)	(41.10)
US entry age ≥18*son		-1.97*			-39.90	
		(1.18)			(68.44)	
Asian immigrant *son			-1.66			-59.48
			(1.43)			(84.65)
Immigrant		-0.79	-0.14		-113.13**	-39.34
		(0.84)	(0.56)		(52.33)	(36.53)
US entry age ≥18	-0.52	-0.21		-77.81**	-17.21	
	(0.53)	(0.69)		(32.96)	(44.04)	
Asian immigrant			-1.66			-193.59**
			(1.43)			(79.70)
R-squared	0.044	0.045		0.052	0.053	0.052
Mothers						
Son	0.35	0.36	0.23	30.40	30.61	24.95
	(0.53)	(0.53)	(0.53)	(22.23)	(22.22)	(22.20)
Immigrant*son	-0.39	-2.38	0.73	-37.06	-149.98*	10.52
	(1.33)	(2.25)	(1.50)	(53.34)	(90.89)	(57.84)
US entry age ≥18*son		2.53			151.87	
		(2.64)			(105.46)	
Asian immigrant *son			-1.43			-39.44
			(2.85)			(111.39)
Immigrant	-9.01***	-0.98	-9.74	-318.50***	3.31	-334.68***
	(1.14)	(1.71)	(1.24)	(46.02)	(71.13)	(48.15)
US entry age ≥18		-12.25***			-495.01***	
		(1.88)			(77.25)	
Asian immigrant			-0.34			-77.24
			(2.58)			(102.06)
R-squared	0.111	0.119	0.110	0.100	0.107	0.102

[a] Sample excludes married couples where one is U.S. born and the other is a first-generation immigrant. Sample sizes are 12,209 fathers and 12,209 mothers.

Notes: Standard errors in parentheses. Survey weights used. Significant levels: *** p<0.01, ** p<0.05, * p<0.1. Control variables include race/ethnicity, age of child, quadratic in age of respondent and spouse, family non-labor income, region, year, education categories of respondent and spouse, and a constant. In the specifications with age of entry, we also control for ambiguous age of entry and its interaction with son.

Table 9. Differential Effects of Child Gender by Race and Ethnicity (1990 Census)

Fathers	*Weeks Last Year*			*Annual Hours*			*Hours Last Week*		
Son	0.21**	0.21**	0.23***	4.53	4.56	6.89	-0.14	-0.14	-0.13
	(0.09)	(0.09)	(0.09)	(6.02)	(6.02)	(6.06)	(0.11)	(0.11)	(0.11)
Black *son	-0.36	-0.36	-0.33	-12.37	-12.61	-9.11	0.47	0.47	0.48
	(0.55)	(0.55)	(0.55)	(29.67)	(29.66)	(29.66)	(0.57)	(0.57)	(0.57)
Asian *son	0.38	0.38	0.74	12.12	12.09	56.91	0.40	0.40	0.52
	(0.61)	(0.61)	(0.70)	(38.89)	(38.89)	(43.99)	(0.67)	(0.67)	(0.77)
Hispanic *son	-0.60	-0.61	-0.41	-41.67	-41.83	-16.64	-0.19	-0.19	-0.12
	(0.45)	(0.45)	(0.48)	(26.94)	(26.95)	(29.21)	(0.46)	(0.46)	(0.50)
Native Amer/Alaskan or	-1.25	-1.29	-1.28	-127.30	-128.83	-127.38	-3.44*	-3.44*	-3.43*
Other race *son	(1.79)	(1.79)	(1.79)	(98.92)	(98.90)	(98.87)	(1.89)	(1.89)	(1.89)
Immigrant*son			-0.46			-57.24**			-0.15
			(0.45)			(27.62)			(0.49)
Immigrant		-0.97***	-0.74**		-38.11***	-8.68		0.12	0.19
		(0.23)	(0.32)		(13.94)	(20.10)		(0.25)	(0.35)
R-squared	0.05	0.05	0.05	0.05	0.05	0.05	0.04	0.04	0.04

Mothers	*Weeks Last Year*			*Annual Hours*			*Hours Last Week*		
Son	0.06	0.06	0.09	5.25	5.40	5.95	0.02	0.02	0.01
	(0.20)	(0.20)	(0.20)	(8.24)	(8.23)	(8.30)	(0.14)	(0.14)	(0.14)
Black *son	-0.08	-0.04	0.00	10.86	12.14	13.14	-0.09	-0.07	-0.08
	(0.85)	(0.85)	(0.85)	(36.59)	(36.60)	(36.61)	(0.68)	(0.68)	(0.68)
Asian *son	-1.15	-1.18	-0.68	-41.08	-41.77	-29.09	0.37	0.34	0.21
	(1.11)	(1.10)	(1.27)	(47.96)	(47.79)	(54.25)	(0.82)	(0.82)	(0.94)
Hispanic *son	-0.94	-0.86	-0.60	-49.81	-47.25	-40.56	-0.51	-0.49	-0.56
	(0.76)	(0.75)	(0.83)	(30.95)	(30.81)	(34.28)	(0.55)	(0.54)	(0.60)
Native Amer/Alaskan or	-2.09	-2.03	-2.01	-104.47	-102.41	-101.84	1.01	1.01	1.00
Other race *son	(2.65)	(2.65)	(2.65)	(125.23)	(124.53)	(124.54)	(1.83)	(1.83)	(1.83)
Immigrant*son			-0.62			-15.72			0.16
			(0.81)			(33.71)			(0.58)
Immigrant	(0.76)	(0.75)	(0.83)		-153.97***	-145.78***		-2.84***	-2.92***
	-2.09	-2.03	-2.01		(17.08)	(24.12)		(0.30)	(0.42)
R-squared	(2.65)	(2.65)	(2.65)	0.06	0.06	0.06	0.06	0.06	0.06

Notes: Sample sizes are 64,442 for weeks and annual hours and 95,365 for hours worked last week. Standard errors are in parentheses. Survey weights are used. Significance levels: ** = 5%; * = 10%. Control variables include race, Hispanic ethnicity, age of child, quadratic in age of respondent and spouse, family non-labor income, region, education categories for respondent and spouse, and a constant.

Table 10. Differential Effects of Child Gender by Race and Ethnicity (2000 Census)

Fathers	Weeks Last Year			Annual Hours			Usual Weekly Hours Last Year		
Son	0.01	0.01	0.02	-10.50	-10.44	-11.55*	-0.22*	-0.22*	-0.24*
	(0.10)	(0.10)	(0.10)	(6.92)	(6.92)	(6.99)	(0.12)	(0.12)	(0.12)
Black *son	-0.29	-0.25	-0.23	1.40	3.36	1.63	-0.46	-0.45	-0.48
	(0.61)	(0.61)	(0.61)	(33.76)	(33.74)	(33.81)	(0.59)	(0.59)	(0.59)
Asian *son	-0.74	-0.76	-0.59	-10.22	-11.52	-27.51	0.12	0.11	-0.23
	(0.52)	(0.52)	(0.64)	(33.48)	(33.45)	(40.31)	(0.58)	(0.58)	(0.70)
Native Amer/Alaskan or	-1.87	-1.93	-1.92	-13.66	-17.22	-18.83	1.02	1.00	0.96
Other race *son	(1.84)	(1.85)	(1.85)	(105.21)	(105.88)	(106.03)	(1.78)	(1.78)	(1.79)
Hispanic *son	0.32	0.34	0.45	33.66	34.91	25.02	0.22	0.23	0.02
	(0.43)	(0.43)	(0.47)	(24.51)	(24.47)	(27.69)	(0.44)	(0.44)	(0.49)
Immigrant*son			-0.21			19.07			0.40
			(0.43)			(26.18)			(0.46)
Immigrant		-1.32***	-1.21***		-73.04***	-82.86***		-0.57**	-0.78**
		(0.22)	(0.31)		(13.42)	(18.86)		(0.24)	(0.34)
R-squared	0.05	0.05	0.05	0.05	0.05	0.05	0.03	0.03	0.03

Mothers	Weeks Last Year			Annual Hours			Usual Weekly Hours Last Year		
Son	0.04	0.04	0.10	-4.50	-4.25	-3.21	-0.07	-0.07	-0.04
	(0.23)	(0.23)	(0.23)	(9.74)	(9.71)	(9.85)	(0.19)	(0.19)	(0.19)
Black *son	-0.45	-0.26	-0.17	23.32	30.46	32.12	0.68	0.83	0.88
	(0.94)	(0.94)	(0.94)	(41.40)	(41.47)	(41.49)	(0.78)	(0.78)	(0.78)
Asian *son	0.83	0.67	1.46	52.74	47.25	62.68	1.23	1.12	1.56
	(0.96)	(0.95)	(1.13)	(40.59)	(40.19)	(47.88)	(0.84)	(0.83)	(0.98)
Native Amer/Alaskan or	-2.61	-2.24	-2.16	-13.90	-0.81	0.82	2.17	2.44	2.49
Other race *son	(2.65)	(2.66)	(2.66)	(113.02)	(113.04)	(113.08)	(2.41)	(2.39)	(2.39)
Hispanic *son	-0.41	-0.37	0.11	-29.86	-28.32	-18.93	-0.89	-0.85	-0.58
	(0.69)	(0.68)	(0.78)	(28.54)	(28.15)	(32.53)	(0.59)	(0.58)	(0.66)
Immigrant*son			-0.95			-18.54			-0.53
			(0.75)			(31.32)			(0.63)
Immigrant		-8.09***	-7.60***		-291.05***	-281.48***		-6.00***	-5.73***
		(0.38)	(0.54)		(15.97)	(22.79)		(0.32)	(0.46)
R-squared	0.07	0.08	0.08	0.06	0.07	0.07	0.05	0.06	0.06

Notes: Sample sizes are 54,120 for weeks, annual hours and usual weekly hours last year. Standard errors are in parentheses. Survey weights are used. Significance levels: ** = 5%; * = 10%. Control variables include race, Hispanic ethnicity, age of child, quadratic in age of respondent and spouse, family non-labor income, region, education categories for respondent and spouse, and a constant.

Table 11. Differential Effects of Child Gender by Race and Ethnicity (1994-2008 CPS)

	Weeks last year			Annual hours			Hours worked last week		
Fathers	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Son	0.25	0.26	0.33	25.60	26.21	28.67*	0.19	0.20	0.25
	(0.23)	(0.23)	(0.23)	(15.99)	(15.99)	(16.16)	(0.33)	(0.33)	(0.34)
Black * son	-0.66	-0.69	-0.57	-61.44	-64.50	-60.32	1.30	1.25	1.36
	(1.22)	(1.22)	(1.22)	(66.19)	(66.24)	(66.56)	(1.52)	(1.52)	(1.52)
Asian * son	-1.96	-1.88	-1.10	-136.00*	-129.01*	-101.53	-0.57	-0.42	0.22
	(1.31)	(1.31)	(1.38)	(70.44)	(70.02)	(77.76)	(1.33)	(1.33)	(1.50)
Other race * son	-2.66	-2.69	-2.69	-330.18***	-332.75***	-332.84***	-6.34**	-6.31**	-6.32**
	(1.97)	(1.97)	(1.98)	(126.53)	(126.37)	(126.36)	(3.15)	(3.15)	(3.14)
Hispanic * son	0.04	0.05	0.63	-38.30	-37.86	-17.50	0.02	0.03	0.50
	(0.62)	(0.62)	(0.74)	(34.83)	(34.85)	(44.67)	(0.77)	(0.77)	(0.96)
Immigrant * son			-1.03			-36.11			-0.84
			(0.72)			(46.41)			(0.97)
Immigrant		-0.96***	-0.43		-85.73***	-67.05*		-2.29***	-1.86**
		(0.37)	(0.55)		(23.62)	(35.01)		(0.50)	(0.73)
R-squared	0.042	0.043	0.043	0.050	0.052	0.052	0.035	0.036	0.036
Mothers									
Son	0.36	0.40	0.33	31.26	32.47	30.92	0.27	0.28	0.16
	(0.55)	(0.55)	(0.55)	(22.89)	(22.85)	(23.19)	(0.39)	(0.39)	(0.40)
Black * son	1.34	0.42	0.28	56.79	22.85	19.36	-0.13	-0.52	-0.76
	(2.19)	(2.18)	(2.18)	(94.47)	(93.65)	(93.70)	(1.69)	(1.68)	(1.69)
Asian * son	-2.65	-2.56	-3.38	-119.45	-116.32	-136.50	-2.56*	-2.50*	-3.99**
	(2.23)	(2.21)	(2.52)	(89.33)	(88.61)	(100.89)	(1.51)	(1.50)	(1.72)
Other race * son	3.92	3.82	3.74	45.87	42.30	40.25	-2.04	-1.67	-1.82
	(4.67)	(4.58)	(4.58)	(202.15)	(199.18)	(199.07)	(3.95)	(3.88)	(3.87)
Hispanic * son	0.14	0.22	-0.36	-10.48	-7.54	-21.85	0.04	0.17	-0.86
	(1.36)	(1.34)	(1.62)	(54.88)	(54.16)	(65.45)	(0.92)	(0.90)	(1.11)
Immigrant * son			1.02			25.17			1.87*
			(1.61)			(64.59)			(1.12)
Immigrant		-8.77***	-9.29***		-322.54***	-335.41***		-5.45***	-6.40***
		(0.83)	(1.17)		(33.20)	(47.87)		(0.58)	(0.82)
R-squared	0.098	0.111	0.111	0.090	0.100	0.100	0.073	0.080	0.080

Notes: Sample sizes are 12,835 for weeks and annual hours and 19,441 for hours worked last week. Standard errors are in parentheses. Survey weights used. Significant levels: *** p<0.01, ** p<0.05, * p<0.1. Control variables include age of child, quadratic in age of respondent and spouse, family non-labor income, region, year, education categories of respondent and spouse, and a constant.

Table 12. Differential Effects of Child Gender by Race and Ethnicity (2003-2008 CPS)

	Weeks last year			Annual hours			Hours worked last week		
Fathers	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
Son	-0.12	-0.11	-0.06	15.07	15.78	18.55	-0.08	-0.10	-0.02
	(0.32)	(0.32)	(0.32)	(21.42)	(21.39)	(21.84)	(0.48)	(0.48)	(0.48)
Black * son	-0.45	-0.44	-0.34	-69.46	-68.75	-63.59	-0.17	-0.12	0.04
	(2.11)	(2.11)	(2.12)	(104.28)	(104.51)	(105.34)	(2.38)	(2.39)	(2.40)
Asian * son	-4.17**	-4.03**	-3.51*	-248.36**	-234.75**	-208.66**	-1.07	-0.77	0.02
	(1.92)	(1.91)	(1.98)	(97.42)	(96.52)	(105.83)	(1.70)	(1.69)	(2.01)
Other race * son	-2.03	-2.06	-2.06	-341.73**	-344.99**	-345.05**	-6.08*	-6.08*	-6.10*
	(2.11)	(2.10)	(2.10)	(149.18)	(147.04)	(146.97)	(3.24)	(3.22)	(3.22)
Hispanic * son	-0.55	-0.61	-0.21	-46.12	-52.06	-32.08	1.04	0.94	1.51
	(0.96)	(0.97)	(1.17)	(52.12)	(52.41)	(63.97)	(1.15)	(1.16)	(1.43)
Immigrant * son			-0.71			-35.39			-1.04
			(1.07)			(60.88)			(1.40)
Immigrant		-1.16**	-0.79		-110.79***	-92.33**		-2.49***	-1.97*
		(0.57)	(0.82)		(31.84)	(42.23)		(0.73)	(1.03)
R-squared	0.041	0.043	0.043	0.052	0.055	0.055	0.038	0.040	0.040
Mothers									
Son	-0.60	-0.54	-0.57	-1.01	1.27	2.28	0.61	0.62	0.57
	(0.78)	(0.77)	(0.79)	(31.89)	(31.81)	(32.40)	(0.55)	(0.55)	(0.56)
Black * son	2.31	1.77	1.71	62.42	42.36	44.42	2.12	1.94	1.86
	(3.56)	(3.51)	(3.53)	(151.00)	(149.95)	(150.29)	(2.41)	(2.36)	(2.38)
Asian * son	0.83	0.96	0.67	15.61	20.57	30.91	-2.50	-2.21	-2.66
	(2.93)	(2.86)	(3.22)	(118.20)	(116.23)	(131.49)	(1.95)	(1.94)	(2.24)
Other race * son	7.92	7.47	7.42	198.03	181.49	183.17	3.52	3.65	3.60
	(5.85)	(5.84)	(5.84)	(260.76)	(262.68)	(263.06)	(4.85)	(4.84)	(4.84)
Hispanic * son	1.25	1.51	1.29	36.44	45.81	53.54	0.08	0.41	0.09
	(2.04)	(2.01)	(2.37)	(80.88)	(79.75)	(95.39)	(1.32)	(1.30)	(1.56)
Immigrant * son			0.40			-13.90			0.60
			(2.19)			(88.74)			(1.50)
Immigrant		-9.59***	-9.80***		-354.59***	-347.26***		-6.43***	-6.74***
		(1.12)	(1.60)		(45.28)	(65.28)		(0.78)	(1.11)
R-squared	0.097	0.113	0.113	0.087	0.101	0.101	0.075	0.085	0.085

Notes: Sample sizes are 6,135 for weeks and annual hours and 9,461 for hours worked last week. Standard errors are in parentheses. Survey weights are used. Significant levels: *** p<0.01, ** p<0.05, * p<0.1. Control variables include age of child, quadratic in age of respondent and spouse, family non-labor income, region, year, education categories of respondent and spouse, and a constant.

Appendix Table 1. Effects of Child Gender for 1st or 2nd Generation Immigrant Parents Families versus Native Parent Families[a] (1994-2008 CPS)

	Weeks worked last year	Annual hours
Fathers		
Son	0.27	20.56
	(0.22)	(15.21)
1st or 2nd generation immigrant*son	-1.41**	-81.45**
	(0.71)	(39.62)
1st or 2nd generation immigrant	-0.23	-86.91**
	(0.56)	(34.25)
R-squared	0.044	0.052
Mothers		
Son	0.26	24.65
	(0.52)	(21.89)
1st or 2nd generation immigrant*son	0.06	-9.56
	(1.42)	(55.91)
1st or 2nd generation immigrant	-9.99***	-369.52***
	(1.19)	(46.92)
R-squared	0.112	0.101

[a] Sample excludes married couples where one is U.S. born and the other is a first-generation immigrant. Sample sizes are 12,247 fathers and 12,247 mothers.

Notes: Standard errors in parentheses. Survey weights used. Significant levels: *** p<0.01, ** p<0.05, * p<0.1. Control variables include race/ethnicity, age of child, quadratic in age of respondent and spouse, family non-labor income, region, year, education categories of respondent and spouse, and a constant.